

# Fighting Enemies and Noise: Competition of Residents and Invaders in a Stochastically Fluctuating Environment

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**Abstract.** The possible control of competitive invasion by infection of the invader and multiplicative noise is studied. The basic model is the Lotka-Volterra competition system with emergent carrying capacities. Several stationary solutions of the non-infected and infected system are identified as well as parameter ranges of bistability. The latter are used for the numerical study of invasion phenomena. The diffusivities, the infection but in particular the white and coloured multiplicative noise are the control parameters. It is shown that not only competition, possible infection and mobilities are important drivers of the invasive dynamics but also the noise and especially its color and the functional response of populations to the emergence of noise.

**Keywords and phrases:** eco-epidemiological model, explicit and emergent carrying capacities, bioinvasion, resident–invader competition, biocontrol, infection, standard incidence, diffusion, multiplicative noise, coloured noise, functional response to noise

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## 1. Introduction

The main aim of modeling biological population dynamics is to improve the understanding of the functioning of food chains and webs as well as their dependence on internal and external conditions. Hence, mathematical models of biological population dynamics have not only to account for growth and interactions but also for spatiotemporal processes like random or directed and joint or relative motion of species, as well as the heterogeneity of the environment. Early attempts began with statistics, exponential growth, physicochemical (neutral) diffusion, and Lotka-Volterra type interactions. These approaches have been continuously refined to more realistic descriptions of the development of natural populations.

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Ecological and epidemiological models are known since more than 200 years. First attempts to merge these models appeared only about 30 years ago, cf. [1, 8, 10, 14] as well as [62, 63]. Infectious diseases are prominent examples of biological invasions and continue to (re-)emerge in modern times. The negative econo-ecological effects of bioinvasions [7, 45] have led to a remarkable hype of bioinvasion research incl. modeling, cf. [18, 41, 48, 50, 66].

The history of research on stochastic processes and integration is long as well, historical surveys have been published, cf. [22, 38]. The seminal work by Itô [21] and Stratonovich [57] should be particularly recognised. Modeling environmental variability with multiplicative white noise goes back to the 1970s. Not only did May [34, 35] introduce the model that is used until today as a perturbation of the growth rate of a population by “white noise” but only a few years later, a more mechanistic basis of this model was developed. Branching processes provide a stochastic model that describes the number of offspring for a given number of individuals  $Z_i$  within one generation  $i$ . The population number  $Z_{i+1}$  in generation  $i+1$  is updated for a given population  $Z_i$  according to previously chosen probability distributions. This model was extended by Smith and Wilkinson [54, 55] by a stochastic process  $\zeta_i$  which modulates for each new generation the offspring that is generated. Branching processes in random environments (BPRE) provide an individual-based model for population growth. For large population numbers, a BPRE can be approximated by a stochastic differential equation (SDE) that accounts both for demographic as well as environmental stochasticity. Keiding [24] conjectured the form of the resulting diffusion approximation, his conjecture was rigorously proven by Kurtz [27]. The model introduced by May is obtained from Kurtz’ diffusion approximation by neglecting the term due to demographic stochasticity. Thus, the SDE model for environmental stochasticity is derived from the influence of a random environment on the population dynamics of a branching process but neglects its demographic stochasticity. Conceptually, it is straightforward to account for spatial heterogeneities by extending a system of SDEs by diffusion terms that represent the spatial spread of the individual species. In the resulting system of reaction-diffusion equations, the stochastic processes  $\zeta_i$  are replaced by time-dependent random fields.

Perturbation with multiplicative white noise is the most commonly used model for stochastic environmental fluctuations. The focus of this study is to generalise this model by relaxing the assumptions that noise is multiplicative and uncorrelated. On the one hand, a multiplicative noise term implies that the effect of environmental fluctuations on the individuals of a population is additive. Whereas this seems reasonable for small population densities we propose that for large population numbers the effect of individual responses to environmental fluctuations on the population should decrease. Thus, we suggest that the population-dependent response to environmental noise saturates for large population numbers similar to the functional response of predators at large prey densities and therefore we model the population-dependent response to environmental noise in a completely analogous way.

On the other hand, the assumption of uncorrelated white noise is usually justified by the coarseness of temporal or spatial scale, respectively. If spatial or temporal correlation length are much shorter than the time or length scale of interest, it seems valid to consider the time-dependent random field as uncorrelated. However, particular care must be taken when considering spatiotemporal dynamics driven by noise. Stochastic differential equations driven by uncorrelated noise can usually be solved over a function space such as  $L^2$  and this remains true if the system is extended to a reaction-diffusion equation over one-dimensional space ( $d = 1$ ). But for spatial dimensions  $d \geq 2$  solutions for stochastic reaction-diffusion equations driven by uncorrelated noise can only be guaranteed in a space of generalised functions, see e.g. [42, 65]. The reason for this phenomenon is that the Laplacian cannot smooth uncorrelated noise sufficiently for spatial dimensions exceeding 1 so that a solution may contain peaks resembling the  $\delta$  distribution. Not only is the physical significance of these solutions debatable but also numerical approximations are not capable of capturing this aspect of the continuous system. Here we take a pragmatic point of view on this difficult problem and present numerical solutions for temporally and spatially white

noise where the space may be interpreted as a discrete lattice whose nodes interact by a discrete Laplacian.

Apart from the mathematical problems arising with white noise terms, the correlation structure of ecological data, in particular time series data, has been extensively studied. The most commonly used class of models consists of power-law noises, the so-called  $1/f^\beta$  noises. Here,  $1/f^\beta$  describes the density of the Fourier-transformed signal over the frequency  $f$ . By analogy with spectra of visible light, some values of the exponent  $\beta$  are associated with colours. For example, pink noise and red noise refer to  $\beta = 1$  and  $\beta = 2$ , respectively, whereas the flat spectrum ( $\beta = 0$ ) is characteristic of white noise. More generally,  $\beta$  is also referred to as the “reddening” of a signal. In his pioneering analysis of environmental noise, Steele combined theoretical models and empirical data and proposed that whereas time series data from terrestrial ecosystem are approximately uncorrelated or white, spectra from marine data are reddened [56]. Twenty years later, Vasseur and Yodzis generally confirmed the work of Steele based on an analysis of various sources of data [61]. At the same time, Halley developed a theoretical argument why  $1/f^\beta$ -noises with  $\beta > 0$  and, in particular, (pink)  $1/f$ -noise, are preferable to uncorrelated white noise as a model for stochastic environmental fluctuations [15, 16]. The influence of temporally coloured noise has been intensively investigated, primarily in the most simple population dynamics models, cf. the review by Ruokolainen et al. [47], whereas Siekmann and Malchow probably published the only study that considered spatially coloured (but temporally white) noise [52, 53]. Although the  $1/f^\beta$  family of noise models has clearly been the most popular model for stochastic environmental fluctuations, it is difficult to rule out alternative stochastic processes. In fact, even Halley—who generally propagates the  $1/f$  model—argues that it is advisable to consider different alternative noise models. For this reason we investigate the spatio-temporal noise model due to García-Ojalvo et al. [12, 13], a spatio-temporal generalisation of the Ornstein-Uhlenbeck process. In contrast to  $1/f^\beta$  noises the density of this noise decays exponentially in Fourier space. A further interesting alternative to the models mentioned so far are bounded noise models, an emerging field of research, discussed, for example, in a recent monograph edited by d’Onofrio [6].

In this study we compare the effects of nonlinear response to noise as well as spatiotemporal correlations to generic multiplicative white noise for the example of a spatial eco-epidemiological model. The model represents the biocontrol of an invasive species through targeted infection [29, 30]. In Section 2 we present the eco-epidemiological model with some stability results (Section 2.1) and the general form of stochastic spatiotemporal models with nonlinear correlated noise terms considered in this study (Section 2.2). Because the numerical solution of stochastic partial differential equations can be challenging, we briefly describe the numerical methods we use for integrating the model considered here including a discussion of the generation of spatiotemporally coloured noise. In Section 4 we demonstrate that for a given parameterisation of our eco-epidemiological invasion model, different parameters of the noise model may lead to qualitatively very different behaviour. Different properties of the stochastic perturbations may not only lead to opposite predictions of survival and extinction but also considerably affect the pattern of the spatiotemporal dynamics. We briefly discuss some implications of our study in Section 5

## 2. Resident-invader competition with infection in the invader population

For the invasion of a resident population by a competing invader, the Lotka-Volterra competition model is used, i.e.,

$$\frac{dN_1}{dt} = r_1 N_1 - c_{11} N_1^2 - c_{12} N_1 N_2, \quad (2.1)$$

$$\frac{dN_2}{dt} = r_2 N_2 - c_{22} N_2^2 - c_{21} N_1 N_2, \quad (2.2)$$

where  $N_1$  and  $N_2$  are resident and invader respectively. Carrying capacities will not explicitly be introduced because they can suppress a higher variety of solutions and rather appear as special cases

[9, 26, 32, 51]. The  $r$ 's stand for the growth rates and the  $c$ 's for the inter- and intraspecific competition.

A specific infection of the invading population can be used as biocontrol measure to stop and reverse the invasion, cf. [5, 17, 23, 37]. To model this, the invader population is split into susceptibles  $S$  and infecteds  $I$ ,

$$N_2 = S + I.$$

Then, the local dynamics reads with notation  $\mathbf{X} = \{X_1 \equiv N_1, X_2 = S, X_3 = I\}$

$$\frac{dX_1}{dt} = f_1(\mathbf{X}) = r_1 X_1 - c_{11} X_1^2 - c_{12} X_1 (X_2 + X_3), \quad (2.3)$$

$$\frac{dX_2}{dt} = f_2(\mathbf{X}) = r_2 X_2 - c_{22} X_2 (X_2 + X_3) - c_{21} X_1 X_2 - \lambda \frac{X_2 X_3}{(X_2 + X_3)^q}, \quad (2.4)$$

$$\frac{dX_3}{dt} = f_3(\mathbf{X}) = r_2 X_3 - c_{22} X_3 (X_2 + X_3) - c_{21} X_1 X_3 + \lambda \frac{X_2 X_3}{(X_2 + X_3)^q} - \mu X_3, \quad (2.5)$$

where  $\lambda$  is the transmission coefficient of the disease and  $\mu$  the disease-induced higher mortality rate of the infecteds. The exponent  $q$  allows to describe mass-action type ( $q = 0$ ) and frequency-dependent transmission ( $q = 1$ ) of the disease respectively [1, 36].

However, one cannot expect that growth rates and competition intensities of susceptibles and infecteds are the same. They should rather be split and could be ordered like

$$\begin{aligned} r_2 &\Rightarrow \{r_2, r_3\} && \& r_3 \leq r_2, \\ c_{12} &\Rightarrow \{c_{12}, c_{13}\} && \& c_{13} \leq c_{12}, \\ c_{21} &\Rightarrow \{c_{21}, c_{31}\} && \& c_{21} \leq c_{31}, \\ c_{22} &\Rightarrow \{c_{22}, c_{23}, c_{32}, c_{33}\} && \& c_{23} \leq c_{33} \leq c_{22} \leq c_{32}. \end{aligned} \quad (2.6)$$

The ordering of the intra- and interspecific competition coefficients of susceptibles and infecteds depends on the biological species, cf. [2]. However, it can be certainly accepted [51] that

$$c_{23} \leq c_{22} \wedge c_{33} \leq c_{32}.$$

System (2.3–2.5) then changes to

$$\frac{dX_1}{dt} = f_1(\mathbf{X}) = (r_1 - c_{11} X_1) X_1 - (c_{12} X_2 + c_{13} X_3) X_1, \quad (2.7)$$

$$\frac{dX_2}{dt} = f_2(\mathbf{X}) = (r_2 - c_{22} X_2) X_2 - (c_{21} X_1 + c_{23} X_3) X_2 - \lambda \frac{X_2 X_3}{(X_2 + X_3)^q}, \quad (2.8)$$

$$\frac{dX_3}{dt} = f_3(\mathbf{X}) = (r_3 - \mu - c_{33} X_3) X_3 - (c_{31} X_1 + c_{32} X_2) X_3 + \lambda \frac{X_2 X_3}{(X_2 + X_3)^q}. \quad (2.9)$$

For convenience, the model of the local dynamics is not analysed in terms of  $X_1$ ,  $X_2$  and  $X_3$  but rather in  $X_1$ ,  $i$  and  $N_2 = X_2 + X_3$  where  $i$  is the prevalence, i.e., the infected fraction of the total invader population  $N_2$  [19],

$$i = \frac{X_2}{X_2 + X_3} = \frac{X_2}{N_2}.$$

Having in mind that

$$\frac{di}{dt} = \frac{1}{N_2} \left( \frac{dX_2}{dt} - i \frac{dN_2}{dt} \right), \quad (2.10)$$

it follows

$$\frac{dX_1}{dt} = r_1 X_1 - c_{11} X_1^2 - [c_{12}(1-i) + c_{13}i] X_1 N_2, \quad (2.11)$$

$$\frac{di}{dt} = \left\{ r_3 - r_2 + \lambda N_2^{1-q} - \mu + N_2 [(c_{22} - c_{32})(1-i) + (c_{23} - c_{33})i] + X_1 [c_{21} - c_{31}] \right\} * i(1-i), \quad (2.12)$$

$$\frac{dN_2}{dt} = \{G(N_2, i) - [c_{21}(1-i) + c_{31}i] X_1\} N_2, \quad (2.13)$$

with

$$G(N_2, i) = r_2(1-i) + (r_3 - \mu)i - N_2 [c_{22}(1-i)^2 + (c_{23} + c_{32})i(1-i) + c_{33}i^2]. \quad (2.14)$$

The latter expression is also found for predator-prey systems with infected prey [51]. Note that if the resident  $X_1$  resp. the predator in [51] cannot distinguish between susceptible and infected invader resp. prey, the temporal change of the prevalence becomes independent of the type of interspecific ecological interactions such as competition and predation. It only contains terms describing the intraspecific competition of susceptibles and infecteds in the infected population.

## 2.1. Stationary solutions and stability for frequency-dependent (standard) incidence $q=1$

In phytopathology, the transmission of especially fungal diseases is described with standard incidence [60]. A corresponding model of the invasion of a fungal disease over a vineyard has been investigated in [4]. Further on, only the standard incidence is considered, i.e.,  $q = 1$ .

The infection-free system, i.e.,  $i \equiv 0, N_2 = X_2$ , is the Lotka-Volterra competition model with its known stationary solutions and their stability ranges. Especially interesting for the consideration of spatial invasions is the bistable parameter range

$$r_1 c_{22} - r_2 c_{12} > 0 \wedge r_2 c_{11} - r_1 c_{21} > 0,$$

when both the invader-free  $(\frac{r_1}{c_{11}}, 0, 0)$  and the resident-free  $(0, 0, \frac{r_2}{c_{22}})$  states are stable and can compete for space. The opposite case may also happen: The invader arrives already infected, i.e.,  $i \equiv 1, N_2 = X_3$ , and the invader-free  $(\frac{r_1}{c_{11}}, 0, 0)$  and the resident-free  $(0, 0, \frac{r_3 - \mu}{c_{33}})$  states can be both at once stable for

$$r_1 c_{33} - (r_3 - \mu) c_{13} > 0 \wedge (r_3 - \mu) c_{11} - r_1 c_{31} > 0.$$

However, the latter as well as the possible bistability of resident-only  $(\frac{r_1}{c_{11}}, 0, 0)$  and (susceptible-infected)-invader-only  $(0, i^{S1}, N_2^{S1})$  states will not be considered here.

## 2.2. Spatiotemporal dynamics in a variable environment

The main focus of this study is to consider the spatiotemporal effects of a more detailed model of environmental variability. We assume that all species spread randomly so that mobility can be described as diffusion with coefficients  $\mathbf{D} = \{D_{ii} = D_i; D_{ij} \equiv 0 \forall i \neq j; i, j = 1, 2, 3\}$ . Also we add Gaussian random fields  $\boldsymbol{\xi}(\mathbf{r}, t) = \{\xi_i(\mathbf{r}, t); i = 1, 2, 3\}$  to system (2.7-2.9) so that we obtain the system of stochastic partial differential equations

$$\frac{\partial \mathbf{X}(\mathbf{r}, t)}{\partial t} - \mathbf{D} \Delta \mathbf{X}(\mathbf{r}, t) = \mathbf{f}[\mathbf{X}(\mathbf{r}, t)] + \mathbf{g}[\mathbf{X}(\mathbf{r}, t)] \boldsymbol{\xi}(\mathbf{r}, t), \quad (2.15)$$

where the matrix function  $\mathbf{g}(\mathbf{X}) = \{g_{ij}(\mathbf{X}); i, j = 1, 2, 3\}$  determines the density-dependent noise intensity. We consider horizontal processes with position vector  $\mathbf{r} = \{x, y\}$  and corresponding Laplace

operator  $\Delta = \partial^2/\partial x^2 + \partial^2/\partial y^2$ . In literature, often temporally and spatially uncorrelated “white” Gaussian fields with zero mean and delta correlation have been considered

$$\langle \xi_i(\mathbf{r}, t) \rangle = 0, \quad \langle \xi_i(\mathbf{r}_1, t_1) \xi_i(\mathbf{r}_2, t_2) \rangle = \delta(\mathbf{r}_1 - \mathbf{r}_2) \delta(t_1 - t_2), \quad i = 1, 2, 3. \quad (2.16)$$

Here, we investigate the effect of extending this model by correlated “coloured” noise with correlation lengths  $\tau$  and  $\lambda$  in the temporal and spatial domain, respectively. Apart from using coloured noise we also investigate a generalisation of the density-dependent noise  $\mathbf{g}(\mathbf{X})$ . Purely diagonal, linear multiplicative noise

$$g_{ii}(\mathbf{X}) = \omega_{ii} X_i; \quad g_{ij}(\mathbf{X}) = 0 \quad \forall i \neq j; \quad i, j = 1, 2, 3; \quad (2.17)$$

can be interpreted as a model where individuals respond independently to stochastic environmental variability. Thus, the effect of environmental fluctuations on each individual directly translates into variability at the population level – the response at the population level is additive. The alternative model suggested here is based on the assumption that in large populations individuals do not respond to fluctuations independently from each other. Instead we propose that larger populations respond to environmental variability in a more robust way, i.e., neither favourable nor adverse effects influence the population proportional to the number of individuals:

$$g_{ij}(\mathbf{X}) = \frac{\omega_{ij} X_j^m}{\gamma_{ij} + \sum_{k=1}^3 a_{ik} X_k^n}; \quad i, j = 1, 2, 3; \quad 1 \leq m \leq n \leq 2. \quad (2.18)$$

For  $m = n$ , the parameter  $\omega_{ij}/a_{ij}$  is the maximum noise intensity that is reached asymptotically for large populations  $X_j$ . The parameter  $\gamma_{ij}$  is the population level at which the noise intensity reaches half of the maximum level  $\omega_{ij}/a_{ij}$ . Thus, this parameter describes the ability of a population to collectively reduce the effect of noise – the higher  $\gamma_{ij}$ , the higher the population level must be until population  $X_j$  is appreciably affected by environmental variability.

For  $m < n$ , the noise intensity even decreases and eventually vanishes for high population densities. However, these values are never reached.

In previous papers [29,30], it was shown that a certain variability of the environment and the mobilities of the competitors are the system-driving forces. Extreme events such as landslides lead to bare ground re-invadable by both resident and alien species. These events at random times, size and locations are not considered here. They are replaced by white and coloured noises [46,52,53]. Again, the biocontrol of the invasion through a specific infection of the non-indigenous species is studied.

### 3. Numerical methods

The numerical solution of stochastic partial differential equations is a difficult problem and the subject of current research. For this reason we explain how the spatiotemporal model (2.15) can be solved numerically and how spatiotemporally correlated noise can be generated. We follow a finite difference approach where in a first step the spatial domain is discretised. In this way the system of stochastic partial differential equations is approximated by uncoupled stochastic differential equations that are solved numerically in a second step.

#### 3.1. Peaceman-Rachford method

For the first step, we use the semi-implicit Peaceman-Rachford method [44, 58] which, in particular for stochastic equations, often seems to be more robust than the simpler explicit schemes. The two-dimensional space  $(x, y)$  is discretised as  $(j\delta x, k\delta y)$  with spatial step sizes  $(\delta x, \delta y)$  respectively. Time  $t$  is split into  $n\delta t$  with time steps  $\delta t$ . Then, the spatiotemporally explicit scheme with centered difference approximation of the second spatial derivative of  $X$  is at time step  $(n+1)$  and inner position  $(j, k)$

$$\frac{X_{jk}^{n+1} - X_{jk}^n}{\delta t} = \frac{X_{j-1k}^n - 2X_{jk}^n + X_{j+1k}^n}{\delta x^2} + \frac{X_{jk-1}^n - 2X_{jk}^n + X_{jk+1}^n}{\delta y^2}. \quad (3.1)$$

Spatial grid points at the system's border have to account for the boundary conditions. As mentioned before, this explicit scheme is not robust enough for highly variable values of  $X$ . Spatially semi-implicit schemes with time splitting and alternating spatial directions are the better choice. One of them is the Peaceman-Rachford method which holds for two-dimensional space. The first half time step, the derivative with respect to  $x$  is computed implicitly and explicitly with respect to  $y$ , and at the second half time step *vice versa*:

$$\frac{X_{jk}^{n+1/2} - X_{jk}^n}{\delta t/2} = \frac{X_{j-1k}^{n+1/2} - 2X_{jk}^{n+1/2} + X_{j+1k}^{n+1/2}}{\delta x^2} + \frac{X_{jk-1}^n - 2X_{jk}^n + X_{jk+1}^n}{\delta y^2}, \quad (3.2)$$

$$\frac{X_{jk}^{n+1} - X_{jk}^{n+1/2}}{\delta t/2} = \frac{X_{j-1k}^{n+1/2} - 2X_{jk}^{n+1/2} + X_{j+1k}^{n+1/2}}{\delta x^2} + \frac{X_{jk-1}^{n+1} - 2X_{jk}^{n+1} + X_{jk+1}^{n+1}}{\delta y^2}. \quad (3.3)$$

Spatially semi-implicit schemes are much easier to handle with lower numerical costs compared to fully implicit ones. It has been proven that the Peaceman-Rachford scheme is unconditionally stable [20].

### 3.2. Derivative-free Milstein method

In the following we will explain how the discretised interaction-noise system can be solved using the derivative-free Milstein method and how spatiotemporally correlated noise can be generated. For numerical integration, the derivative-free Milstein method is used, cf. [25, 39, 40] but also the short descriptions [11, 49]. Sometimes and in particular for the purpose of this study, it is even sufficient to consider purely diagonal intensity matrices

$$g_{ii}(\mathbf{X}) = \frac{\omega_{ii} X_i^m}{\gamma_{ii} + a_{ii} X_i^n}; \quad a_{ii} \equiv 1; \quad i = 1, 2, 3; \quad 1 \leq m \leq n \leq 2. \quad (3.4)$$

Then, the Milstein scheme reads with time step  $\Delta t$  and Stratonovich interpretation

$$X_{t+\Delta t}^i = X_t^i + f_i(X_t^i)\Delta t + g_{ii}(X_t^i)\Delta W_i + \frac{1}{2\sqrt{\Delta t}} [g_{ii}(\bar{X}_t^i) - g_{ii}(X_t^i)] (\Delta W_i)^2, \quad (3.5)$$

with

$$\bar{X}_t^i = X_t^i + f_i(X_t^i)\Delta t + g_{ii}(X_t^i)\sqrt{\Delta t},$$

and

$$\Delta W_i = W_{t+\Delta t}^i - W_t^i \sim \sqrt{\Delta t} \mathcal{N}(0, 1).$$

As usual,  $\mathcal{N}(0, 1)$  stands for the normal distribution with zero mean and unity variance. The required uniformly distributed random numbers are generated with the Mersenne Twister [33], the normally distributed with the common Box-Muller algorithm [3].

### 3.3. Generation of correlated Gaussian random fields

García-Ojalvo and Sancho [12, 13] developed a method for generating spatially and temporally coloured noise from the stochastic reaction-diffusion equation

$$\frac{\partial \zeta(\mathbf{r}, t)}{\partial t} = \frac{\lambda^2}{\tau} \Delta \zeta(\mathbf{r}, t) - \frac{1}{\tau} \zeta(\mathbf{r}, t) + \frac{1}{\tau} \eta(\mathbf{r}, t) \quad (3.6)$$

Here, the term  $\eta(\mathbf{r}, t)$  stands for uncorrelated (white) noise. The parameters  $\tau$  and  $\lambda$  determine the correlation lengths in the temporal and the spatial domain, respectively. In addition, a scaling factor  $\epsilon$  for the variance of normally-distributed random variables in Fourier space has to be chosen. The authors introduce Eq. (3.6) as an analogon of the Ornstein-Uhlenbeck process which is the solution of (3.6) without the spatial term  $\frac{\lambda^2}{\tau} \Delta \zeta(\mathbf{r}, t)$ . But it has to be noted that for two-dimensional space (and spatial dimensions

exceeding two) the solutions of equation (3.6) will be generalised functions rather than functions in a space such as  $L^2$ . Thus, strictly speaking, we are generating discrete spatiotemporally coloured random fields that are derived from the model (3.6) without being approximate solutions of the continuous problem.

The spatiotemporal random field  $\zeta(\mathbf{r}, t)$  is simulated by transforming a discretised version  $\zeta_{ij}(t)$  of (3.6) to Fourier space:

$$\zeta_{\mu\nu}(t) = (\Delta x)^2 \sum_{i,j=1}^L \exp(-\mathbf{i}k \cdot \mathbf{r}) \zeta_{ij}(t) \quad (3.7)$$

We denote the discrete Fourier transform  $\zeta_{\mu\nu}(t)$  with greek indices  $\mu, \nu$  rather than  $i, j$  and the coordinate  $k$  in frequency space is

$$k = \frac{2\pi}{L\Delta x}(\mu, \nu), \quad \mu, \nu = 0, \dots, L-1.$$

Then the Fourier transformed field  $\zeta_{\mu\nu}(t)$  at the next time step  $t + \Delta t$  is calculated by

$$\zeta_{\mu\nu}(t + \Delta t) = \zeta_{\mu\nu}(t) \exp\left(-\frac{c_{\mu\nu}}{\tau} \Delta t\right) + \sqrt{\frac{\epsilon(L\Delta x)^2}{\tau c_{\mu\nu}} \left[1 - \exp\left(-\frac{c_{\mu\nu}}{\tau} \Delta t\right)\right]} \alpha_{\mu\nu} \quad (3.8)$$

Here,  $\alpha_{\mu\nu}$  is the Fourier transform of an uncorrelated Gaussian noise field  $\alpha_{ij}$ . The efficiency of this method is increased by directly generating the Fourier transformed field  $\alpha_{\mu\nu}$ . The (complex-valued) discrete Fourier transform  $\alpha_{\mu\nu}$  of real-valued fields  $\alpha_{ij}$  obeys some symmetries that lead to the following restrictions:

$$\alpha_{\mu\nu} = \alpha_{L-\mu, L-\nu}^* \quad (3.9)$$

$$\alpha_{\mu\nu} \in \mathbb{R}, \quad \text{for } \mu, \nu = 0, \frac{L}{2} \quad (3.10)$$

where  $z^*$  denotes the complex conjugate of  $z \in \mathbb{C}$ . The condition (3.9) means that  $\alpha_{\mu\nu}$  have the same real part as  $\alpha_{L-\mu, L-\nu}$  found by reflecting through the centre ( $\mu = \nu = L/2$ ) whereas the imaginary parts only differ by opposite signs.

Also in (3.8),  $c_{\mu\nu}$  is the Fourier transform of the discretisation of the differential operator  $\mathcal{L} = \text{id} - \lambda^2 \Delta$ :

$$c_{\mu\nu} = 1 - \frac{2\lambda^2}{(\Delta x)^2} \left[ \cos\left(\frac{2\pi\mu}{L}\right) + \cos\left(\frac{2\pi\nu}{L}\right) - 2 \right]. \quad (3.11)$$

For further details on the derivation of this method, we refer the reader to [12] or [13].

## 4. Resident-invader competition-diffusion model with infected invader in a variable environment

### 4.1. Local dynamics with multiplicative white noise and induced transitions

Not surprising and like for  $q = 0$  [59, 64], Lehmann [28] and Woyzichovski [67] found disease-induced oscillations for  $q = 1$  as well. Their interesting result was that there may be bistability of the resident-only state and the limit cycle when coexisting resident, susceptible and infected invaders oscillate, cf. Figure 1. The following parameters have been used for the latter setting:

$$\begin{aligned} r_1 &= 1.2500, \quad r_2 = 1.0000, \quad r_3 = 0.6775 \\ c_{11} &= 0.5000, \quad c_{12} = 0.8500, \quad c_{21} = 0.4250, \quad c_{13} = 0.5000, \quad c_{31} = 0.5000 \\ c_{22} &= 0.6000, \quad c_{23} = 0.4000, \quad c_{32} = 0.6010, \quad c_{33} = 0.5000 \\ \lambda &= 0.4000, \quad \mu = 0.0250 \end{aligned} \quad (4.1)$$



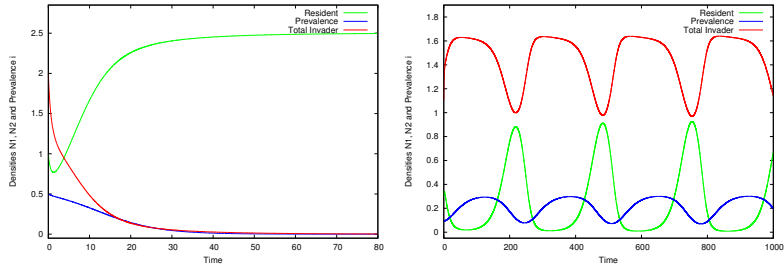
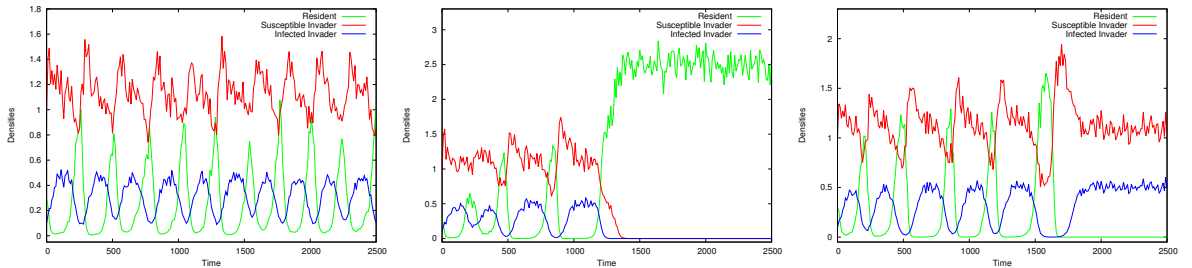


FIGURE 1. Bistability of resident-only state and oscillating coexistence

All other semi-trivial states turn out to be unstable for this parameter range.

Now, one effect of white noise in locally multiple stable systems is shown, i.e., the switch from one stable attractor to the other for sufficiently but not too high noise intensity. For simplicity, the linear density dependence (2.17) of the intensity is chosen, which has been successfully applied to numerous cases. Here, the leaving of the initial limit cycle is demonstrated.

FIGURE 2. Noise-induced transitions from oscillating coexistence to resident-only state resp. resident extinction. Linear density dependence (2.17) with  $\omega_{ii} = 0.1$   $i = 1, 2, 3$ .

The three subfigures of Figure 2 show typical outcomes of hundreds of simulations with different seeds of the random number generator. The left subfigure shows the persistence of the limit cycle whereas the middle demonstrates the expected leaving of the cycle for the other stable stationary solution, i.e., the resident-only state. The result in the right subfigure appeared a bit unexpected, however, due to some catastrophic shift, the resident died out and the remaining susceptible and infected invader survived. One should have in mind that the latter  $(0, X_2^s, X_3^s)$  state is unstable to the reintroduction of the resident.

## 4.2. Invasions and noise I

### 4.2.1. Linear noise and biocontrol of invasion

For the beginning, the results in the mentioned previous papers [29, 30], where simulated landslides led to bare land competitively re-invadable by resident and invader, are reproduced with external noise. The initial condition is a “red” invader patch at its emergent carrying capacity  $r_{22}/c_{22}$  at the “upper left corner” of the “green” habitat of the native species at  $r_{11}/c_{11}$ . This patch should exceed the related critical patch size. Otherwise it will simply decay regardless of its competitive strength and mobility [31, 43]. Zero-flux boundary conditions are applied.

The linear noise (2.17) and parameters from the previous publications have been taken:

$$\begin{aligned}
 r_1 &= 1.000, \quad r_2 = 1.000, \quad r_3 = 0.800 \\
 c_{11} &= 1.000, \quad c_{12} = 1.300, \quad c_{21} = 1.200, \quad c_{13} = 1.299, \quad c_{31} = 1.201 \\
 c_{22} &= 0.999, \quad c_{23} = 0.998, \quad c_{32} = 1.001, \quad c_{33} = 1.000 \\
 \lambda &= 0.405, \quad \mu = 0.200 \\
 \omega_{ii} &= 0.250; \quad D_i = 45.000; \quad i = 1, 2, 3.
 \end{aligned} \tag{4.2}$$

The invader patch spreads and seems to grow unstoppable.

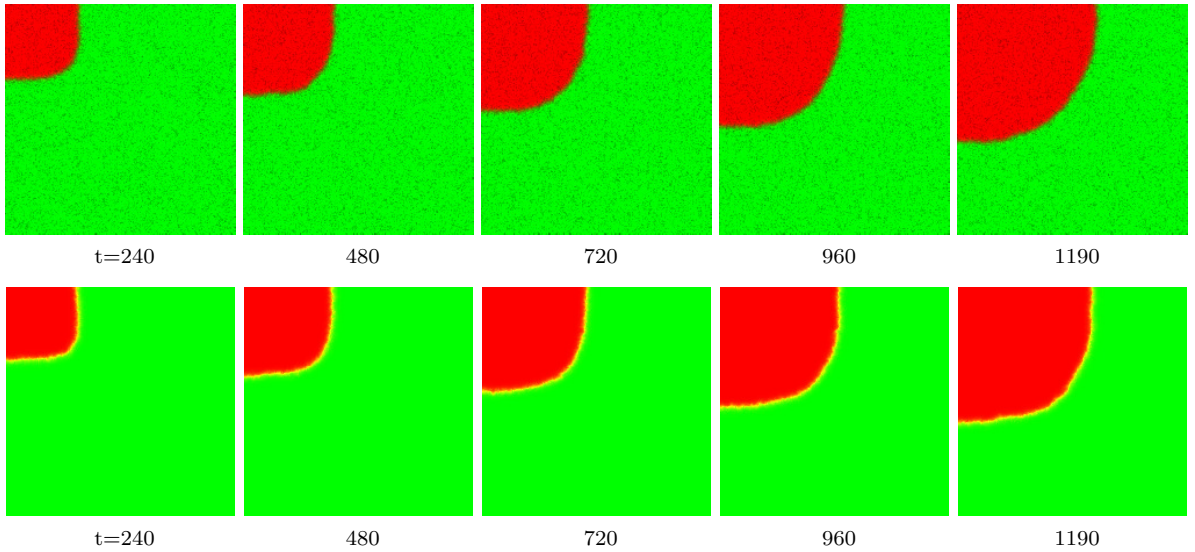


FIGURE 3. Simulation 1: Parameters as given in (4.2). The first row shows results with white noise, the second row shows spatiotemporally coloured noise with  $\omega_{ii} = 0.03$ ,  $\epsilon = 0.001$  and correlation lengths  $\tau = 1$  and  $\lambda = 1$ .

Then, a biological control measure is applied. The invader population is partly infected and the invasion successfully rolled back.

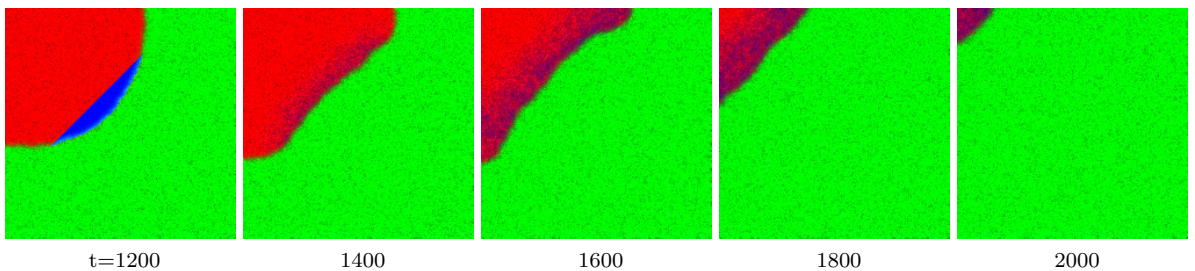


FIGURE 4. Simulation 1 continued: Partial “blue” infection of the invader at  $t=1200$ .

The changes of the fraction of the invaded area can be seen in Figure 5.

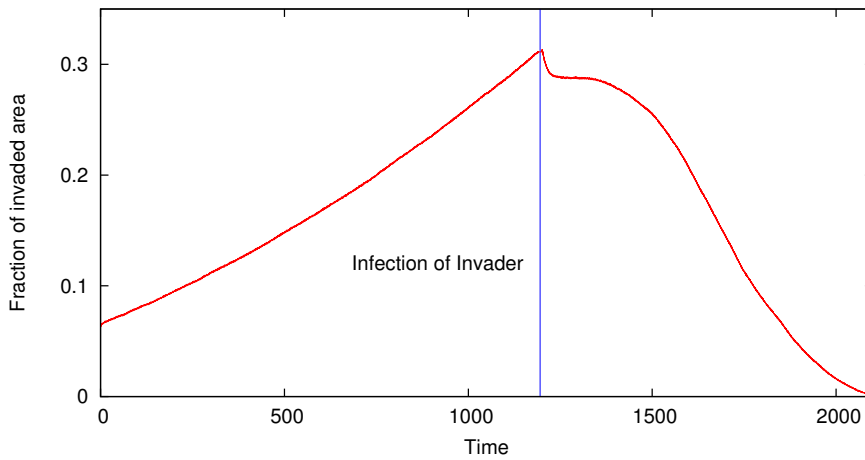


FIGURE 5. Fraction of invaded area before and after partial infection.

#### 4.2.2. Nonlinear response to noise and noise control of invasion

We now consider the saturating response to noise described above (2.18) with

$$\begin{aligned} m = n = 2; \quad \omega_{11} = 50, \quad \omega_{22} = 0.1, \quad \omega_{33} = 10; \\ \gamma_{11} = 100, \quad \gamma_{22} = 3, \quad \gamma_{33} = 300. \end{aligned} \quad (4.3)$$

The extinction of the invader due to hostile environmental conditions is shown in Figure 6. The resident is used and adapted to the environment and happily survives.

This dynamics is only due to the specific noise response of the populations. All growth and interaction parameters remained the same as in sec. 4.2.1

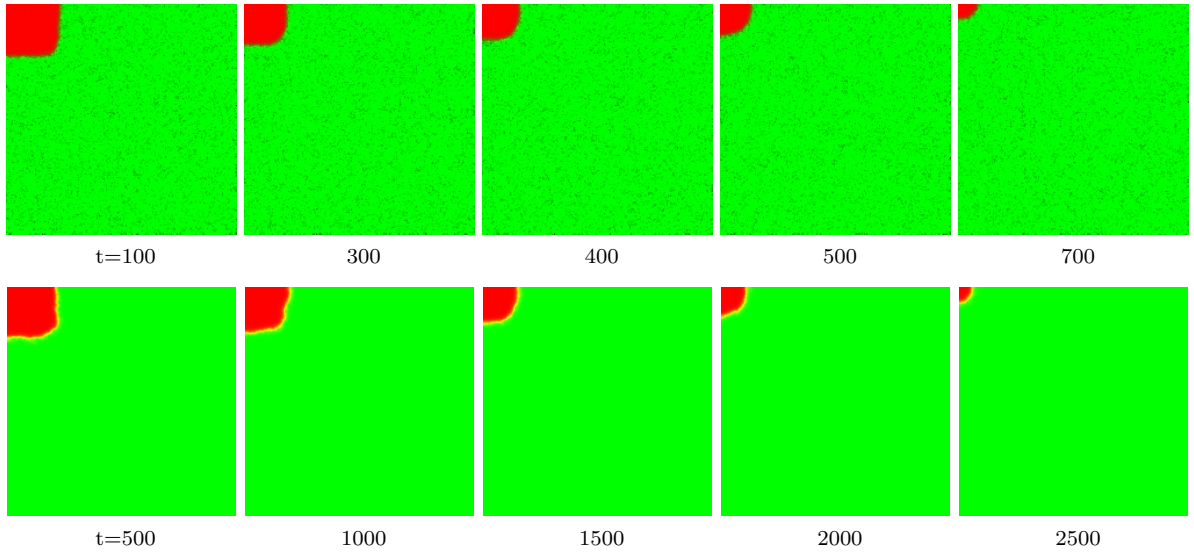


FIGURE 6. Repressed invasion through population-specific nonlinearly density-dependent noise intensities for white noise in the first row. In the second row for coloured noise with correlation lengths  $\tau = 1$  and  $\lambda = 1$ ,  $\omega_{11} = 0.05$ ,  $\omega_{22} = \omega_{33} = 0.01$  qualitatively similar behaviour is observed, albeit on a much slower timescale. However, other simulations have shown that stronger temporal and spatial correlations can support invasions.

### 4.3. Invasions and noise IIa

Coming back to the parameter range of bistability of resident-only state and oscillating coexistence of all three populations, i.e., parameters (4.1), the initial condition is chosen as uniformly populated by the resident  $X_1$  at its emergent carrying capacity. At a defined location at the boundary, an initial patch of the invading populations  $X_2$  and  $X_3$  attempts to spread. Again, simply linear noise (2.17) is applied and the influence of increasing noise intensity  $\omega_{ii} = \omega$ ;  $i = 1, 2, 3$ ; studied.

It is seen in Figure 7 that a successful invasion requires a certain supercritical noise intensity.

Somehow, the resident supports the invasion of its own area. Diffusion and noise enhance the mixing of resident and invaders at the front. Therefore, all three together jump on the stable limit cycle of coexistence and invade the remaining invader-free area, cf. Figure 8.

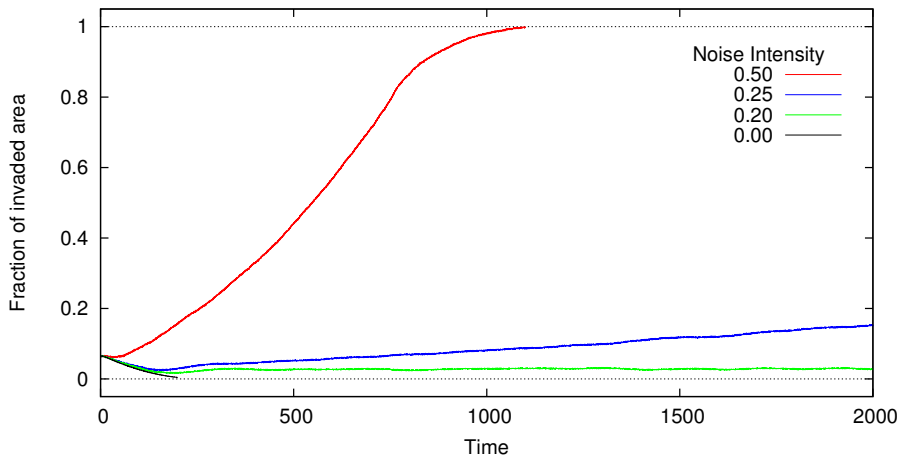


FIGURE 7. Decline resp. growth of invaded area for increasing white noise intensity.

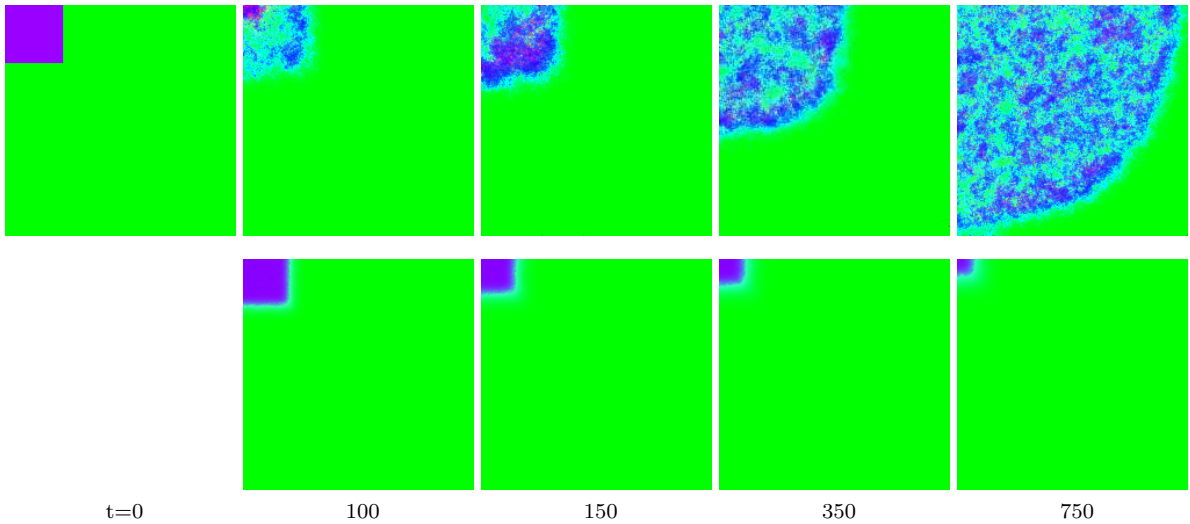


FIGURE 8. Upper row: Successful fast invasion of the resident’s area for linear white noise (2.17) with  $\omega_{ii} = 0.5$ ;  $i = 1, 2, 3$ . Lower row: Stopped invasion of the resident’s area for nonlinear noise (4.3) with  $\omega_{ii} = 0.25$ ;  $i = 1, 2, 3$ ;  $\gamma_{11} = 1.00$ ,  $\gamma_{22} = 25.00$ ,  $\gamma_{33} = 4.00$ .

The following Figure 9 shows that weakly correlated noise still allows for invasion but stronger correlated does not.

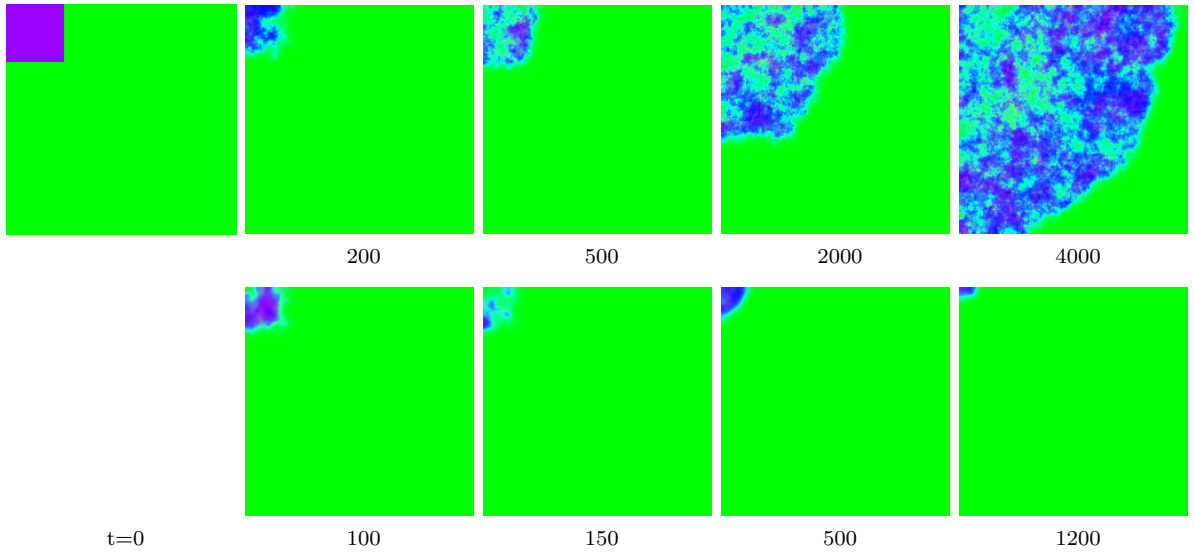


FIGURE 9. Same parameters as in Fig. 8. The invasion is successful for coloured noise and correlation lengths  $\tau = \lambda = 1$  but unsuccessful for  $\tau = \lambda = 15$ . Other parameters  $\epsilon = 0.001$ ,  $\omega_{ii} = 0.05$ .

After some difficulties at the beginning that can be the end for the invader at lower noise intensities, the purple invader patch turns into the blue of the limit cycle. Finally, the resident survives but has to share its habitat with the aliens. However, if the nonlinear response or coloured noise is applied, the invasion can be stopped and rolled back again.

#### 4.4. Invasions and noise IIb

Now, it is assumed that the limit cycle of resident and invaders has already invaded most of the area and only a small part is left for the resident alone. Again, the parameters (4.1) are used. For this situation several interesting patterns appear that are again purely due to different properties of the environmental noise.

##### 4.4.1. Dependence on noise intensity

It is surprising that the resident turns out to be strong enough to defeat the invasion as far the noise is below a subcritical threshold. The fraction of invaded area over time is plotted in Figure 10. An example for the defeat of invasion is given in Figure 11.

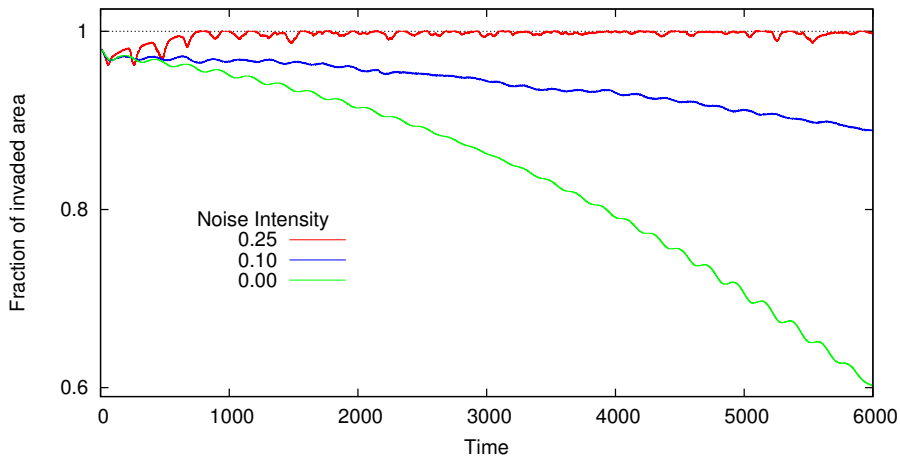


FIGURE 10. Decline resp. growth of invaded area for increasing white noise intensity.

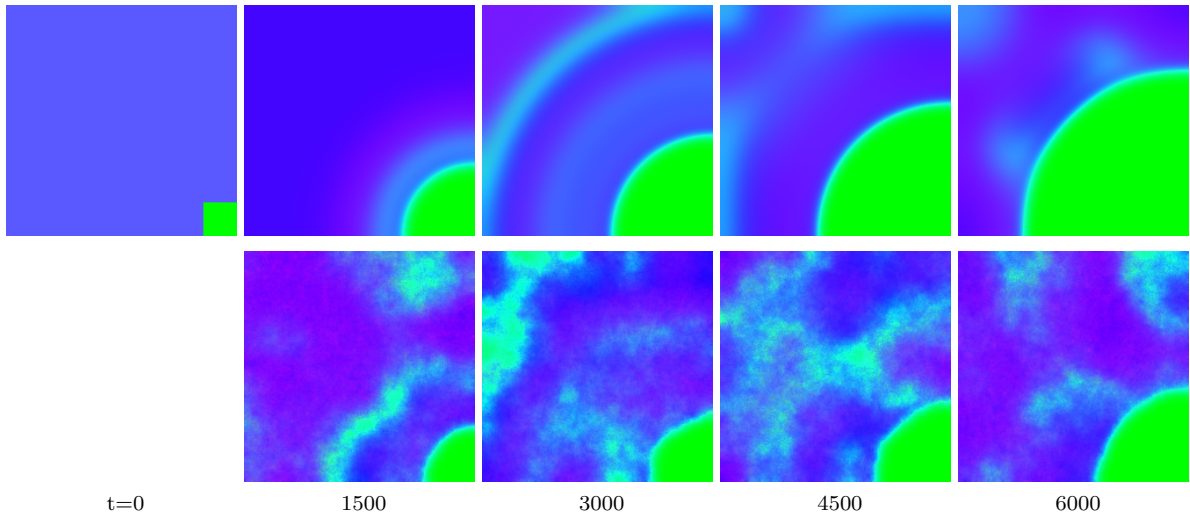


FIGURE 11. Finally successful defeat of invasion for no noise (upper row) resp. white noise with  $\omega_{ii} = 0.1$ ;  $i = 1, 2, 3$ .

A preliminary conclusion is that increasing linearly density-dependent white noise supports invasion. In Figure 12, the cloudy result for  $\omega_{ii} = 0.25$  is shown.

Increasing the noise intensity and both correlation lengths ( $\tau = \lambda = 20$ ) leads to a situation where native and resident gain and lose control over parts of the spatial domain in an alternating fashion (Figure 13). It seems that the resident is slowly getting the upper hand: at  $t = 7500$  roughly two thirds of the domain are occupied by natives, at  $t = 10000$  the resident has lost a few areas and displaced the invader in a few others.

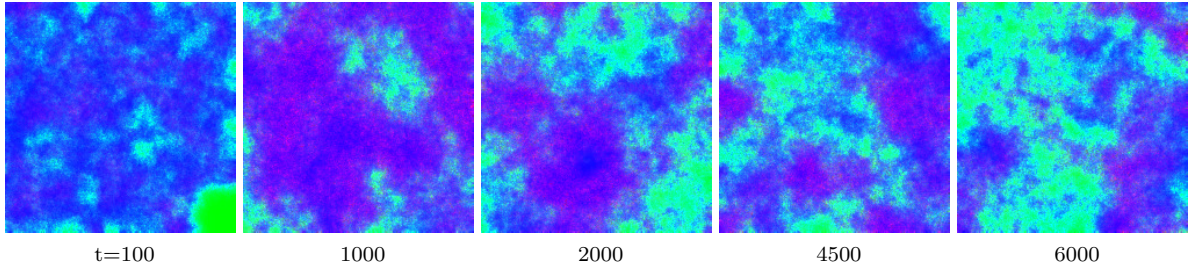


FIGURE 12. Cloudy invasion for linear white noise with  $\omega_{ii} = 0.25$ ;  $i = 1, 2, 3$ . Initial condition as in Figure 11.

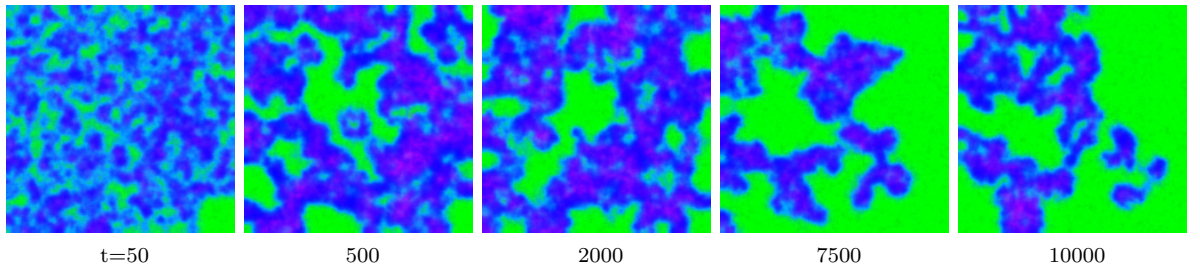


FIGURE 13. Resident and invader fight under coloured noise for territory creating a highly irregular pattern for  $\omega_{ii} = 0.1$ ,  $\epsilon = 0.001$ ,  $\tau = 20$ ,  $\lambda = 20$ . Initial condition as in Figure 11.

#### 4.4.2. Metapopulation patches

This does not necessarily change for nonlinear noise, however, one setting is found where the fraction of invaded area is dropped down from initially 98% to 21%. The resident population splits into three metapopulations, spatially separated by the invader populations, cf. Figure 14.

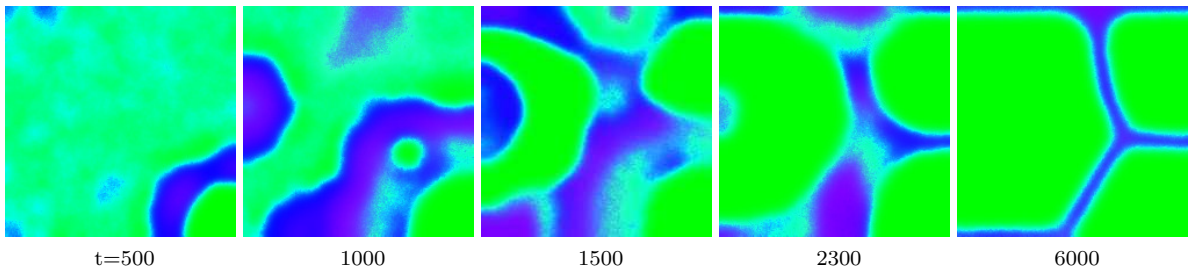


FIGURE 14. Stable formation of spatially separated patches of the resident population for white noise intensity  $\omega_{ii} = 0.25$ ;  $i = 1, 2, 3$ ;  $\gamma_{11} = 1.0$ ,  $\gamma_{22} = 25.0$ ,  $\gamma_{33} = 4.0$ . Initial condition as in Figure 11.

As observed above, higher noise intensities  $\omega$  help the invader to establish by stabilising the coexistence limit cycle. In contrast, stronger correlations i.e. increasing correlation lengths  $\tau$  in time or  $\lambda$  in space generally enable the native species to displace the invaders. Only for relatively high correlation



lengths of 10 or above the native species is able to form patches and avoid a cloudy invasion as in Figure 12.

In Figure 15, for a low noise intensity  $\omega_{ii} = 0.03$ ,  $\tau = 1$  and  $\lambda = 15$  we observe the emergence of a quasi-stationary pattern similar to Figure 14.

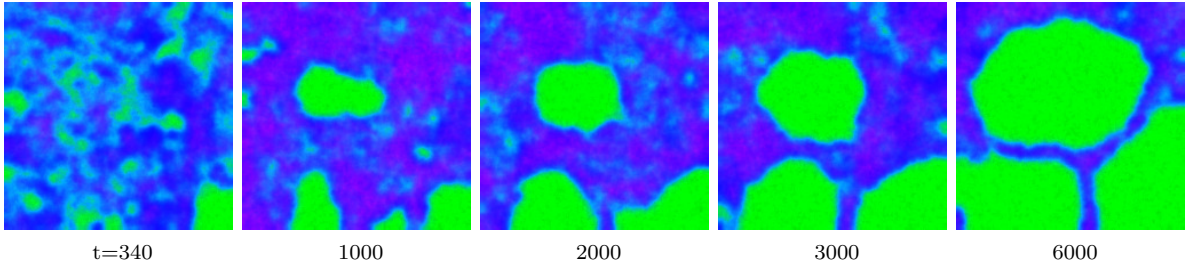


FIGURE 15. Resident slowly displaces invader while forming seemingly stable spatially separated patches, similar to the results in Figure 14 but without nonlinearity in the noise term. Parameters:  $\omega_{ii} = 0.03$ ,  $\epsilon = 0.001$ ,  $\tau = 1$ ,  $\lambda = 15$ . Initial condition as in Figure 11.

#### 4.4.3. Spiral waves

Wavy structures are found as well, however, at the cost of full invasion. One example is presented in Figure 16.

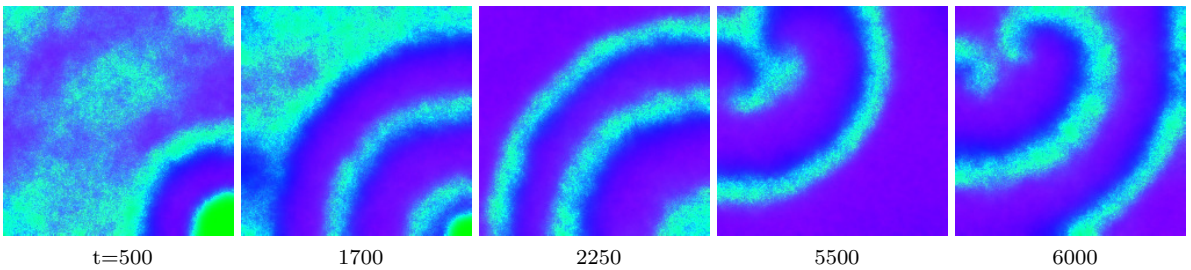


FIGURE 16. Stable formation of spatially separated patches of the resident population for linear white noise intensity  $\omega_{ii} = 0.3$ ;  $i = 1, 2, 3$ ;  $\gamma_{11} = 0.5625$ ,  $\gamma_{22} = 4.0$ ,  $\gamma_{33} = 9.0$ . Initial condition as in Figure 11.

For nonlinear noise, the spiral waves seen in Figure 16 can also be observed if the noise is coloured. For small values of spatial and temporal correlation lengths ( $\tau = \lambda = 1$ ), smaller and slightly more irregular spiral waves can be observed, cf. Figure 17.

For increased correlation in time or space (e.g.  $\tau = 1$  as above but  $\lambda = 5$ ), the native population can still maintain a spreading front that eventually displaces the spiral waves formed in its wake, cf. Figure 18.

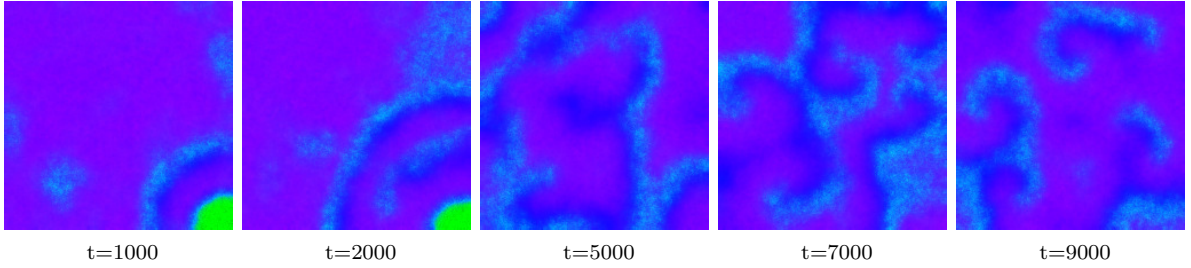


FIGURE 17. Formation of spiral waves – similar to Figure 16 but less regular – for coloured noise intensity  $\omega_{ii} = 0.03$ ;  $i = 1, 2, 3$ ;  $\epsilon = 0.001$ ,  $\tau = 1$ ,  $\lambda = 1$ ,  $\gamma_{11} = 0.5625$ ,  $\gamma_{22} = 4.0$ ,  $\gamma_{33} = 9.0$ . Initial condition as in Figure 11.

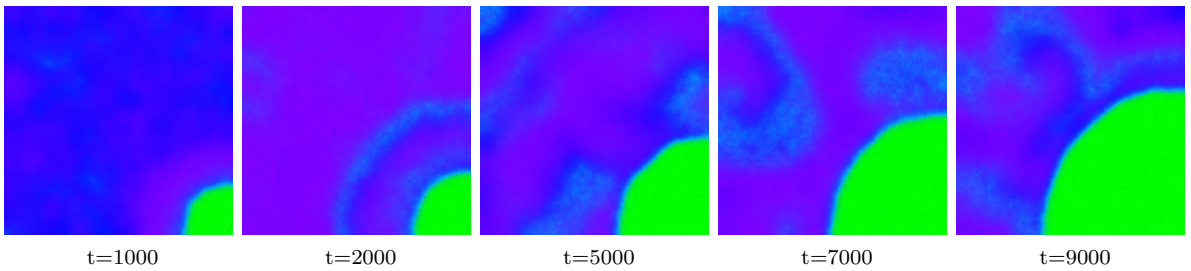


FIGURE 18. For increased spatial correlation the spiral waves are displaced by a front of natives. Parameters:  $\omega_{ii} = 0.03$ ;  $i = 1, 2, 3$ ;  $\epsilon = 0.001$ ,  $\tau = 1$ ,  $\lambda = 5$ ,  $\gamma_{11} = 0.5625$ ,  $\gamma_{22} = 4.0$ ,  $\gamma_{33} = 9.0$ . Initial condition as in Figure 11.

## 5. Conclusions

Populations are exposed to fluctuations of many environmental parameters such as nutrient availability, temperature etc. that may have positive or adverse effects. Whereas it would be impractical to explicitly consider a host of factors that each on their own may only have a small influence on growth or decline of the population it is possible to represent the collective effect of these factors as stochastic environmental variability. The standard model for stochastic environmental variability in population dynamics are stochastic differential equations with a multiplicative noise term.

In this model, the matrix of maximum noise intensities  $\omega$  is the only parameter that can be used for capturing all aspects of environmental fluctuations. The standard assumption that environmental noise is temporally and spatially uncorrelated neglects the fact that many environmental factors are, in fact, typically correlated.

Whereas at first glance it seems that this can be convincingly justified by assuming that the spatial and temporal correlation lengths  $\tau$  and  $\lambda$  of the noise are much shorter than the spatiotemporal scale under consideration, we have demonstrated in this study that using correlated instead of uncorrelated noise may lead to qualitatively very different model behaviour. Thus, neglecting possible correlations may, in fact, lead to different explanations of the observed system behaviour.

Another implicit assumption that has previously been unquestioned is the linear increase of noise intensity with population number. This model implies that environmental effects on each individual in a population simply add up to an overall effect on the population. In our opinion it is likely that

for increasing population numbers, perturbations should not independently affect each individual but rather saturate due to interactions of the individuals so that the collective response of the population saturates to a maximum noise intensity for large population numbers. This model requires an additional matrix  $\gamma$  which characterises the ability of the population to “buffer” stochastic fluctuations: for low values of  $\gamma_{ii}$ , population  $X_i$  is exposed to intensities close to the maximum noise level  $\omega_{ii}$  for low or moderate population numbers whereas for a population with a large  $\gamma_{ii}$  the noise intensity  $\omega_{ii}$  is only reached for high population numbers.

The results presented in this paper have been obtained for purely diagonal noise intensity matrices. More complex forms are left to future work.

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